

# CS 58500 – Theoretical Computer Science Toolkit

Lecture 5 (02/03)

Concentration Inequality IV

[https://ruizhezhang.com/course\\_spring\\_2026.html](https://ruizhezhang.com/course_spring_2026.html)



# Today's Lecture

- Subgaussian and Subgamma Random Variable
- The Entropy Method
- Talagrand's Inequality
- Applications: LIS and TSP

# Subgaussian Random Variable

A random variable  $X$  is called  $\sigma^2$ -subgaussian if its log-MGF satisfies

$$\psi(\theta) := \log \mathbb{E}[\exp(\theta(X - \mathbb{E}[X]))] \leq \theta^2 \sigma^2 / 2 \quad \forall \theta \in \mathbb{R}$$

We call  $\sigma^2$  the variance proxy.

Equivalently,

- $\Pr[|X - \mathbb{E}[X]| \geq t] \leq 2e^{-t^2/2\sigma^2}$
- $\mathbb{E}[|X - \mathbb{E}[X]|^k] \leq \sigma^k k^{k/2}$  for any  $k \in \mathbb{Z}_+$

Examples:

- True Gaussian random variable  $\mathcal{N}(0, \sigma^2)$
- Bounded random variable: if  $a \leq X \leq b$  a.s., then  $X$  is  $(b - a)^2/4$ -subgaussian

# Subgaussian Random Variable

A random variable  $X$  is called  $\sigma^2$ -subgaussian if its log-MGF satisfies

$$\psi(\theta) := \log \mathbb{E}[\exp(\theta(X - \mathbb{E}[X]))] \leq \theta^2 \sigma^2 / 2 \quad \forall \theta \in \mathbb{R}$$

We call  $\sigma^2$  the variance proxy.

Equivalently,

- $\Pr[|X - \mathbb{E}[X]| \geq t] \leq 2e^{-t^2/2\sigma^2}$
- $\mathbb{E}[|X - \mathbb{E}[X]|^k] \leq \sigma^k k^{k/2}$  for any  $k \in \mathbb{Z}_+$

**Lemma.** If  $X_1, X_2$  are independent subgaussian random variables with variance proxy  $\sigma_1^2$  and  $\sigma_2^2$ , then  $X_1 + X_2$  is  $(\sigma_1^2 + \sigma_2^2)$ -subgaussian

➤ It immediately recovers the Hoeffding's inequality

# Subgamma Random Variable

A random variable  $X$  is called  $(\sigma^2, c)$ -subgamma if

$$\psi(\theta) \leq \frac{\theta^2 \sigma^2}{2(1 - |\theta|c)} \leq \frac{\theta^2 \sigma^2}{2} \quad \forall |\theta| < 1/c$$

It holds that

$$\Pr[|X - \mathbb{E}[X]| \geq t] \leq 2 \max \left\{ e^{-\frac{t^2}{2\sigma^2}}, e^{-\frac{t}{2c}} \right\}$$

## Examples:

- If  $X \sim \mathcal{N}(0,1)$ , then  $X^2$  is  $(4,3)$ -subgamma
- If  $X$  is  $\sigma^2$ -subgaussian, then  $X$  is  $(\sigma^2, 0)$ -subgamma
- Bounded random variable: if  $|X - \mathbb{E}[X]| \leq b$  a.s., then  $X$  is  $(\text{Var}[X], b/3)$ -subgamma
- If  $X$  is  $(\sigma^2, c)$ -subgamma, then  $\alpha X$  is  $(\alpha^2 \sigma^2, \alpha c)$ -subgamma

# Subgamma Random Variable

A random variable  $X$  is called  $(\sigma^2, c)$ -subgamma if

$$\psi(\theta) \leq \frac{\theta^2 \sigma^2}{2(1 - |\theta|c)} \leq \frac{\theta^2 \sigma^2}{2} \quad \forall |\theta| < 1/c$$

It holds that

$$\Pr[|X - \mathbb{E}[X]| \geq t] \leq 2 \max \left\{ e^{-\frac{t^2}{2\sigma^2}}, e^{-\frac{t}{2c}} \right\}$$

**Lemma.** If  $X_1, X_2$  are independent subgamma random variables with parameters  $(\sigma_1^2, c_1)$  and  $(\sigma_2^2, c_2)$ , then  $X_1 + X_2$  is  $(\sigma_1^2 + \sigma_2^2, \max\{c_1, c_2\})$ -subgamma

➤ It immediately recovers the Bernstein inequality

# Today's Lecture

- Subgaussian and Subgamma Random Variable
- The Entropy Method
- Talagrand's Inequality
- Applications: LIS and TSP

# The Entropy Method

The entropy of a random variable  $X$  is defined as

$$\text{Ent}[X] := \mathbb{E}[X \log X] - \mathbb{E}[X] \log \mathbb{E}[X]$$

**Lemma (Herbst).** Suppose that

$$\text{Ent}[e^{\theta X}] \leq \frac{\theta^2 \sigma^2}{2} \mathbb{E}[e^{\theta X}] \quad \forall \theta \geq 0$$

Then,  $X$  is  $\sigma^2$ -subgaussian.

## Tensorization of entropy

- For a function  $f(x_1, \dots, x_n)$ , and for each  $i \in [n]$ , define
$$\text{Ent}_i[f(x)] := \text{Ent}[f(x_1, \dots, x_{i-1}, X_i, x_{i+1}, \dots, x_n)]$$
- For independent random variables  $X_1, \dots, X_n$ , we have

$$\text{Ent}[f(X_1, \dots, X_n)] \leq \mathbb{E} \left[ \sum_{i=1}^n \text{Ent}_i[f(X_1, \dots, X_n)] \right]$$



# Proof of Herbst Lemma

*Proof.*

- We'll verify that  $\psi(\theta) := \log \mathbb{E}[\exp(\theta(X - \mathbb{E}[X]))] \leq \theta^2 \sigma^2 / 2$
- $\psi(\theta) = \log \mathbb{E}[e^{\theta X}] - \theta \mathbb{E}[X]$

$$\frac{d}{d\theta} \left( \frac{\psi(\theta)}{\theta} \right) = \frac{\mathbb{E}[X e^{\theta X}]}{\theta \mathbb{E}[e^{\theta X}]} - \frac{\mathbb{E}[X]}{\theta} - \frac{\log \mathbb{E}[e^{\theta X}]}{\theta^2} + \frac{\mathbb{E}[X]}{\theta} = \frac{\mathbb{E}[X e^{\theta X}]}{\theta \mathbb{E}[e^{\theta X}]} - \frac{\log \mathbb{E}[e^{\theta X}]}{\theta^2}$$

- $\text{Ent}[e^{\theta X}] = \theta \mathbb{E}[X e^{\theta X}] - \mathbb{E}[e^{\theta X}] \log \mathbb{E}[e^{\theta X}]$
- Thus,  $\frac{d}{d\theta} \left( \frac{\psi(\theta)}{\theta} \right) = \frac{\text{Ent}[e^{\theta X}]}{\theta^2 \mathbb{E}[e^{\theta X}]} \leq \frac{\sigma^2}{2}$  by assumption
- Then, we have

$$\frac{\psi(\theta)}{\theta} = \int_0^\theta \frac{\text{Ent}[e^{\tau X}]}{\tau^2 \mathbb{E}[e^{\tau X}]} d\tau \leq \int_0^\theta \frac{\sigma^2}{2} d\tau = \frac{\theta \sigma^2}{2}$$



# The Entropy Method

**Lemma** (Discrete Modified log-Sobolev (MLS) Inequality). Let  $f: \mathbb{R} \rightarrow \mathbb{R}$  and let

$$D^-f(x) := f(x) - \inf_y f(y)$$

Then for any random variable  $X$ ,

$$\text{Ent}[e^{f(X)}] \leq \text{Cov}[f(X), e^{f(X)}] \leq \mathbb{E}[|D^-f(X)|^2 e^{f(X)}],$$

where  $\text{Cov}[X, Y] = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y]$

*Proof.*

- For the first inequality,

$$\text{Ent}[e^f] = \mathbb{E}[f e^f] - \mathbb{E}[e^f] \log \mathbb{E}[e^f] \leq \mathbb{E}[f e^f] - \mathbb{E}[e^f] \mathbb{E}[f] = \text{Cov}[f, e^f]$$

- For the second inequality,

$$\text{Cov}[f, e^f] = \mathbb{E}[(f - \inf f)(e^f - \mathbb{E}[e^f])] \leq \mathbb{E}[(f - \inf f)(e^f - e^{\inf f})]$$

- The convexity of  $e^x$  implies that  $e^f - e^{\inf f} \leq e^f (f - \inf f)$



# The Entropy Method: Sharper Bounded Differences

Define one-sided differences for multivariate function:

$$D_i^- f(x) := f(x_1, \dots, x_n) - \inf_z f(x_1, \dots, x_{i-1}, z, x_{i+1}, \dots, x_n)$$

$$D_i^+ f(x) := \sup_z f(x_1, \dots, x_{i-1}, z, x_{i+1}, \dots, x_n) - f(x_1, \dots, x_n)$$

**Theorem** (Bounded differences inequality).

Let  $X_1, \dots, X_n$  be independent random variables. Then,  $f(X_1, \dots, X_n)$  is subgaussian with variance proxy  $2\|\sum_{i=1}^n |D_i f|^2\|_\infty$ . Moreover,

$$\Pr[f - \mathbb{E}[f] \geq t] \leq \exp\left(-\frac{t^2}{4\left\|\sum_{i=1}^n |D_i^- f|^2\right\|_\infty}\right)$$
$$\Pr[f - \mathbb{E}[f] \leq -t] \leq \exp\left(-\frac{t^2}{4\left\|\sum_{i=1}^n |D_i^+ f|^2\right\|_\infty}\right)$$

# The Entropy Method: Sharper Bounded Difference

**Theorem** (Bounded differences inequality).

Let  $X_1, \dots, X_n$  be independent random variables. Then,  $f(X_1, \dots, X_n)$  is subgaussian with variance proxy  $2\|\sum_{i=1}^n |D_i f|^2\|_\infty$

**Theorem** (McDiarmid inequality).

Let  $X_1, \dots, X_n$  be independent random variables. Then,  $f(X_1, \dots, X_n)$  is subgaussian with variance proxy  $\frac{1}{4} \sum_{i=1}^n \|D_i f\|_\infty^2$

In many cases,  $\|\sum_{i=1}^n |D_i f|^2\|_\infty$  can be much smaller than  $\sum_{i=1}^n \|D_i f\|_\infty^2$

# The Entropy Method: Sharper Bounded Difference

**Theorem** (Bounded difference inequality).

Let  $X_1, \dots, X_n$  be independent random variables. Then,  $f(X_1, \dots, X_n)$  is subgaussian with variance proxy  $2\|\sum_{i=1}^n |D_i f|^2\|_\infty$

*Proof.*

- By the [discrete MLS lemma](#),

$$\text{Ent}_i[e^f] \leq \mathbb{E}[|D_i^- f|^2 e^f | X_1, \dots, X_{i-1}, X_{i+1}, \dots, X_n]$$

- By [tensorization](#), for any  $\theta \geq 0$ ,

$$\text{Ent}[e^{\theta f}] \leq \mathbb{E}\left[\sum_{i=1}^n \text{Ent}_i[e^{\theta f}]\right] \leq \mathbb{E}\left[\left(\sum_{i=1}^n |D_i^-(\theta f)|^2\right) e^{\theta f}\right] \leq \theta^2 \left\|\sum_{i=1}^n |D_i^- f|^2\right\|_\infty \mathbb{E}[e^{\theta f}]$$

- We finish the proof by [Herbst lemma](#).



# Today's Lecture

- Subgaussian and Subgamma Random Variable
- The Entropy Method
- Talagrand's Inequality
- Applications: LIS and TSP

# Talagrand's Inequality: Motivating Question

Let  $V$  be a fixed  $d$ -dimensional subspace. Let  $\mathbf{x} \sim \text{Unif}\{-1,1\}^n$ . How well is  $\text{dist}(\mathbf{x}, V)$  concentrated?

- Let  $P$  be the orthogonal projection onto  $V^\perp$ . Then,  $\text{tr}[P] = \dim(V^\perp) = n - d$
- $\text{dist}(\mathbf{x}, V)^2 = |P\mathbf{x} \cdot P\mathbf{x}| = |\mathbf{x}^\top P\mathbf{x}| = \sum_{i,j \in [n]} x_i x_j P_{ij}$
- Thus,  $\mathbb{E}[\text{dist}(\mathbf{x}, V)^2] = \sum_{i \in [n]} P_{ii} = n - d$

How well is  $\text{dist}(\mathbf{x}, V)$  concentrated around  $\sqrt{n - d}$ ?

- Consider  $f(\mathbf{x}) := \text{dist}(\mathbf{x}, V)$  for  $\mathbf{x} \in \{-1,1\}^n$
- For any  $i \in [n]$ , by triangle inequality,  
$$|D_i f(\mathbf{x})| = |\text{dist}(\mathbf{x}_{-i}, V) - \text{dist}(\mathbf{x}, V)| \leq \|\mathbf{x} - \mathbf{x}_{-i}\|_2 = 2$$
- By the **bounded differences inequality**,  $\Pr[|\text{dist}(\mathbf{x}, V) - \sqrt{n - d}| \geq t] \leq 2e^{-2t^2/n}$
- **Useless** since  $\text{dist}(\mathbf{x}, V) \leq \text{dist}(\mathbf{x}, \mathbf{0}) = \sqrt{n}$

# Talagrand's Inequality: Motivating Question

Let  $V$  be a fixed  $d$ -dimensional subspace. Let  $\mathbf{x} \sim \text{Unif}\{-1,1\}^n$ . How well is  $\text{dist}(\mathbf{x}, V)$  concentrated?

- Let  $P$  be the orthogonal projection onto  $V^\perp$ . Then,  $\text{tr}[P] = \dim(V^\perp) = n - d$
- $\text{dist}(\mathbf{x}, V)^2 = |P\mathbf{x} \cdot P\mathbf{x}| = |\mathbf{x}^\top P\mathbf{x}| = \sum_{i,j \in [n]} x_i x_j P_{ij}$
- Thus,  $\mathbb{E}[\text{dist}(\mathbf{x}, V)^2] = \sum_{i \in [n]} P_{ii} = n - d$

How well is  $\text{dist}(\mathbf{x}, V)$  concentrated around  $\sqrt{n - d}$ ?

**Corollary (Talagrand's inequality).** For  $\mathbf{x} \sim \text{Unif}\{-1,1\}^n$ , we have

$$\Pr[|\text{dist}(\mathbf{x}, V) - \sqrt{n - d}| \geq t] \leq C e^{-ct^2}$$

where  $C, c$  are universal constants



# Talagrand's Inequality: Convex Lipschitz Functions

## Theorem (Talagrand).

Let  $A \subseteq \mathbb{R}^n$  be a convex set. Let  $\mathbf{x} \sim \text{Unif}\{0,1\}^n$ . Then

$$\Pr[\mathbf{x} \in A] \Pr[\text{dist}(\mathbf{x}, A) \geq t] \leq e^{-t^2/4} \quad \forall t \geq 0$$

Equivalently, for a convex **1-Lipschitz function**  $f: \mathbb{R}^n \rightarrow \mathbb{R}$  (i.e.,  $|f(x) - f(y)| \leq \|x - y\|_2$  for any  $x, y \in \mathbb{R}^n$ ) and  $\mathbf{x} \sim \text{Unif}\{0,1\}^n$ ,

$$\Pr[f(\mathbf{x}) \leq r] \Pr[f(\mathbf{x}) \geq r + t] \leq e^{-t^2/4} \quad \forall r \in \mathbb{R}, t \geq 0$$

*Proof of the equivalence.*

- “ $\Rightarrow$ ”: let  $A := \{x \in \mathbb{R}^n : f(x) \leq r\}$ . Then  $f$  is convex implies that  $A$  is convex. We also have  $\text{dist}(x, A) \leq t \Rightarrow f(x) \leq r + t$  by the 1-Lipschitzness. Thus,  $\Pr[f(\mathbf{x}) \leq r] = \Pr[\mathbf{x} \in A]$  and  $\Pr[f(\mathbf{x}) \geq r + t] \leq \Pr[\text{dist}(\mathbf{x}, A) \geq t]$
- “ $\Leftarrow$ ”: let  $r = 0$  and  $f(x) = \text{dist}(x, A)$

# Talagrand's Inequality: Convex Lipschitz Functions

## Theorem (Talagrand).

For a convex **1-Lipschitz** function  $f: \mathbb{R}^n \rightarrow \mathbb{R}$  (i.e.,  $|f(x) - f(y)| \leq \|x - y\|_2$  for any  $x, y \in \mathbb{R}^n$ ) and  $\mathbf{x} \sim \text{Unif}\{0,1\}^n$ ,

$$\Pr[f(\mathbf{x}) \leq r] \Pr[f(\mathbf{x}) \geq r + t] \leq e^{-t^2/4} \quad \forall r \in \mathbb{R}, t \geq 0$$

**Corollary.** Let  $\text{med}(X)$  be the **median** of the random variable  $X$ . That is,  $\Pr[X \geq \text{med}(X)] \geq 1/2$  and  $\Pr[X \leq \text{med}(X)] \geq 1/2$ . Then, for a convex 1-Lipschitz function  $f: \mathbb{R}^n \rightarrow \mathbb{R}$  and  $\mathbf{x} \sim \text{Unif}\{0,1\}^n$ ,

$$\Pr[|f(\mathbf{x}) - \text{med}(f(\mathbf{x}))| \geq t] \leq 4e^{-t^2/4}$$

*Proof.*

- Let  $r := \text{med}(f(\mathbf{x}))$ . Then  $\Pr[f(\mathbf{x}) \leq r] \geq 1/2$  and  $\Pr[f(\mathbf{x}) \geq r + t] \leq 2e^{-t^2/4}$
- Let  $r := \text{med}(f(\mathbf{x})) - t$ . Then  $\Pr[f(\mathbf{x}) \geq r + t] \geq 1/2$   $\Pr[f(\mathbf{x}) \leq r] \leq 2e^{-t^2/4}$



# Talagrand's Inequality: Convex Distance

Let the probability space be  $\Omega = \Omega_1 \times \cdots \times \Omega_n$  with product probability measure

## Weighted Hamming distance

- Given  $\alpha \in \mathbb{R}_{\geq 0}^n$ ,  $x, y \in \Omega$ , define

$$d_\alpha(x, y) := \sum_{i=1}^n \alpha_i \mathbf{1}[x_i \neq y_i]$$

- For a subset  $A \subseteq \Omega$ ,  $d_\alpha(x, A) := \inf_{y \in A} d_\alpha(x, y)$

## Talagrand's convex distance

- For  $x \in \Omega$  and  $A \subseteq \Omega$ ,

$$d_T(x, A) := \sup_{\substack{\alpha \in \mathbb{R}_{\geq 0}^n \\ \|\alpha\|_2 = 1}} d_\alpha(x, A)$$

# Talagrand's Inequality: Convex Distance

Let the probability space be  $\Omega = \Omega_1 \times \cdots \times \Omega_n$  with product probability measure

Talagrand's convex distance

$$d_T(x, A) := \sup_{\substack{\alpha \in \mathbb{R}_{\geq 0}^n \\ \|\alpha\|_2 = 1}} \inf_{y \in A} \sum_{i=1}^n \alpha_i \mathbf{1}[x_i \neq y_i]$$

Properties:

- If  $A \subseteq \{0,1\}^n$  and  $x \in \{0,1\}^n$ , then  $d_T(x, A) = \text{dist}(x, \text{conv}(A))$
- For any  $x \in \Omega$ , define  $\phi_x(y) := (\mathbf{1}[x_1 \neq y_1], \dots, \mathbf{1}[x_n \neq y_n]) \in \{0,1\}^n$ , and  $\phi_x(A) := \{\phi_x(y) : y \in A\} \subseteq \{0,1\}^n$  for any  $A \subseteq \Omega$ . Then

$$d_T(x, A) = \text{dist}(\mathbf{0}, \text{conv}(\phi_x(A)))$$

# Talagrand's Inequality: Convex Distance

**Theorem** (Talagrand's inequality, general form).

Let  $A \subseteq \Omega = \Omega_1 \times \cdots \times \Omega_n$ , and  $\mathbf{x} \sim \Omega$  be chosen randomly with independent coordinates. Then

$$\Pr[\mathbf{x} \in A] \Pr[d_T(\mathbf{x}, A) \geq t] \leq e^{-t^2/4}$$

# Talagrand's Inequality: Convex Distance

**Theorem** (Talagrand's inequality, functions with weighted certificates).

Let  $\mathbf{x} \sim \Omega$  with independent coordinates. Suppose that

$$f(y) \geq f(x) - \sum_{i=1}^n \alpha_i(x) \mathbf{1}[x_i \neq y_i] \quad \forall x, y \in \Omega$$

Then,

$$\Pr[|f(\mathbf{x}) - \text{med}(f(\mathbf{x}))| \geq t] \leq 4e^{-t^2/v^2}, \quad v := 2 \sup_{x \in \Omega} \|\alpha(x)\|_2$$

# Talagrand's Inequality: Convex Distance

**Theorem** (Talagrand's inequality, functions with weighted certificates).

Let  $\mathbf{x} \sim \Omega$  with independent coordinates. Suppose that  $f(y) \geq f(x) - \sum_{i=1}^n \alpha_i(x) \mathbf{1}[x_i \neq y_i]$  for any  $x, y \in \Omega$ . Then, for  $v^2 := 4 \sup_{x \in \Omega} \|\alpha(x)\|_2^2$ ,

$$\Pr[|f(x) - \text{med}(f(x))| \geq t] \leq 4e^{-t^2/v^2}$$

*Proof.*

- For  $r \in \mathbb{R}$ , let  $A := \{y : f(y) \leq r - t\}$
- For any  $x \in \Omega$  such that  $f(x) \geq r$ , the assumption gives that
$$\exists \alpha(x) \in \mathbb{R}_{\geq 0}^n, \forall y \in A, \quad d_{\alpha(x)}(x, y) \geq f(x) - f(y) \geq r - (r - t) = t$$
- Then, we have  $d_{\alpha(x)}(x, A) \geq t$  and  $d_T(x, A) \geq t/\|\alpha(x)\|_2 \geq 2t/v$

# Talagrand's Inequality: Convex Distance

**Theorem** (Talagrand's inequality, functions with weighted certificates).

Let  $\mathbf{x} \sim \Omega$  with independent coordinates. Suppose that  $f(y) \geq f(x) - \sum_{i=1}^n \alpha_i(x) \mathbf{1}[x_i \neq y_i]$  for any  $x, y \in \Omega$ . Then, for  $v^2 := 4 \sup_{x \in \Omega} \|\alpha(x)\|_2^2$ ,

$$\Pr[|f(x) - \text{med}(f(x))| \geq t] \leq 4e^{-t^2/v^2},$$

*Proof.*

- Then, we have  $d_{\alpha(x)}(x, A) \geq t$  and  $d_T(x, A) \geq t/\|\alpha(x)\|_2 \geq 2t/v$
- By Talagrand's inequality (general form),

$$\Pr[f \leq r - t] \Pr[f \geq r] \leq \Pr[\mathbf{x} \in A] \Pr\left[d_T(\mathbf{x}, A) \geq \frac{2t}{v}\right] \leq e^{-t^2/v^2}$$

- The result then follows by taking  $r := \text{med}(f) + t$  and  $r := \text{med}(f)$





# Median vs. Mean

For any real random variable  $X$  satisfying

$$\Pr[|X - m| \geq t] \leq 2e^{-t^2/\sigma^2} \quad \forall t \geq 0$$

for some  $m \in \mathbb{R}$  and  $\sigma > 0$ , then the followings hold:

1.  $|\text{med}(X) - m| \leq C\sigma$

2.  $|\mathbb{E}[X] - m| \leq C\sigma$

3. For every constant  $A$ , if  $|m' - m| \leq A\sigma$ , then

$$\Pr[|X - m'| \geq t] \leq 2e^{-\Omega(t^2/\sigma^2)} \quad \forall t \geq 0$$

# Median vs. Mean (Proof)

- We can rescale  $X$  so that  $\sigma = 1$

- For the median, take  $t > \sqrt{2 \log 2}$ :

$$\Pr[|X - m| > \sqrt{2 \log 2}] \leq e^{-t^2} < 1/2$$

Thus,  $\text{med}(X)$  is within  $\sqrt{2 \log 2}$  of  $m$

- For the mean,

$$|\mathbb{E}[X] - m| \leq \mathbb{E}[|X - m|] = \int_0^\infty \Pr[|X - m| \geq t] dt \leq 2 \int_0^\infty e^{-t^2} dt = \sqrt{\pi}$$

- For the last inequality, since  $A$  is constant, by choosing a sufficiently small  $c > 0$ , we can let  $2e^{-ct^2} \geq 1$  when  $t \leq 2A$  (e.g.,  $c = \frac{1}{10A^2}$ ). Then, for  $t > 2A$ , we have

$$\Pr[|X - m'| \geq t] \leq \Pr[|X - m| \geq t/2] \leq e^{-t^2/4}$$



# Today's Lecture

- Subgaussian and Subgamma Random Variable
- The Entropy Method
- Talagrand's Inequality
- Applications: LIS and TSP

# Application 1: Longest Increasing Subsequence

An **increasing subsequence** of a permutation  $\sigma = (\sigma_1, \dots, \sigma_n)$  is defined to be some  $\sigma_{i_1} < \dots < \sigma_{i_\ell}$  for some  $i_1 < \dots < i_\ell$ .

How well is the length  $X$  of the longest increasing subsequence of uniformly random permutation concentrated?

**Example:**  $\sigma = (2, 1, 8, 7, 4, 5, 6, 3)$

- You can show that  $\mathbb{E}[X] = \Theta(\sqrt{n})$  (A good exercise of binomial coefficient approximation)
- For concentration, there is one problem:  $\sigma_1, \dots, \sigma_n$  are **not** independent
- You can sample  $X_1, \dots, X_n \sim_{iid} \text{Unif}[0,1]$ , and their ordering gives a random permutation
- Talagrand's inequality  $\Rightarrow X = \Theta(\sqrt{n}) \pm \mathcal{O}(n^{1/4})$  w.h.p.

# Application 1: Longest Increasing Subsequence

Let  $\Omega = \Omega_1 \times \cdots \times \Omega_n$  and  $A \subseteq \Omega$ . We say  $A$  is **s-certifiable** if for every  $x \in A$ , there exists a subset  $I \subset [n]$  with  $|I| \leq s$  such that for every  $y \in \Omega$ , if  $y_I = x_I$ , then  $y \in A$

- For LIS,  $\Omega = [0,1]^n$  and  $A = \{x \in [0,1]^n : \text{LIS}(x) \geq k\}$ . Then  $A$  is  $k$ -certifiable since we just need to check  $k$  coordinates to determine an increasing subsequence of length  $k$

**Theorem** (Talagrand's inequality for certifiable functions).

Let  $x \sim \Omega$  with independent coordinates. Let  $f: \Omega \rightarrow \mathbb{R}$  be 1-Lipschitz with respect to Hamming distance on  $\Omega$ . Suppose that  $\{x \in \Omega : f(x) \geq r\}$  is  $r$ -certifiable. Then, for  $m = \text{med}(f(x))$ ,

$$\Pr[f(x) \leq m - t] \leq 2e^{-t^2/(4m)}$$

$$\Pr[f(x) \geq m + t] \leq 2e^{-t^2/(4(m+t))}$$

- For  $x \in [0,1]^n$ , let  $f(x) := \text{LIS}(x)$ . Then  $f$  is 1-Lipschitz (since changing one coordinate can change the LIS by at most 1). It is easy to show that  $m = \Theta(\sqrt{n})$ . The above theorem implies the desired concentration bound.

## Theorem (Talagrand's inequality for certifiable functions).

Let  $\mathbf{x} \sim \Omega$  with independent coordinates. Let  $f: \Omega \rightarrow \mathbb{R}$  be 1-Lipschitz with respect to Hamming distance on  $\Omega$ . Suppose that  $\{x \in \Omega : f(x) \geq r\}$  is  $r$ -certifiable. Then, for  $m = \text{med}(f(x))$ ,

$$\Pr[f(\mathbf{x}) \leq m - t] \leq 2e^{-t^2/(4m)}$$

$$\Pr[f(\mathbf{x}) \geq m + t] \leq 2e^{-t^2/(4(m+t))}$$

*Proof.*

- Let  $A := \{f \leq r - t\}$  and  $B := \{f \geq r\}$ . We first show that  $\Pr[f \leq r - t] \Pr[f \geq r] \leq e^{-t^2/(4r)}$
- $B$  is  $r$ -certifiable, so for every  $y \in B$ , let  $I(y)$  denote its certificate with  $|I(y)| \leq r$
- By the 1-Lipschitzness of  $f$ , for every  $x \in A$ ,  $t \leq |f(x) - f(y)| \leq d_H(x, y)$
- We want to apply Talagrand's inequality (the general form):

$$\Pr[\mathbf{x} \in A] \Pr[d_{\mathcal{T}}(\mathbf{x}, A) \geq t] \leq e^{-t^2/4}$$

- For  $i \in [n]$ , define  $\alpha_i(y) := 1/\sqrt{|I(y)|}$  for  $i \in I(y)$  and  $\alpha_i(y) := 0$  otherwise. Then,

$$\|\alpha(y)\|_2 = 1, \quad d_{\alpha}(x, y) \geq t/\sqrt{|I(y)|} \geq t/\sqrt{r} \quad \forall x \in A$$

## Theorem (Talagrand's inequality for certifiable functions).

Let  $\mathbf{x} \sim \Omega$  with independent coordinates. Let  $f: \Omega \rightarrow \mathbb{R}$  be 1-Lipschitz with respect to Hamming distance on  $\Omega$ . Suppose that  $\{x \in \Omega : f(x) \geq r\}$  is  $r$ -certifiable. Then, for  $m = \text{med}(f(x))$ ,

$$\Pr[f(\mathbf{x}) \leq m - t] \leq 2e^{-t^2/(4m)}$$

$$\Pr[f(\mathbf{x}) \geq m + t] \leq 2e^{-t^2/(4(m+t))}$$

*Proof.*

- Let  $A := \{f \leq r - t\}$  and  $B := \{f \geq r\}$ . We first show that  $\Pr[f \leq r - t] \Pr[f \geq r] \leq e^{-t^2/(4r)}$
- We want to apply Talagrand's inequality (the general form):

$$\Pr[\mathbf{x} \in A] \Pr[d_T(\mathbf{x}, A) \geq t] \leq e^{-t^2/4}$$

- For  $i \in [n]$ , define  $\alpha_i(y) := 1/\sqrt{|I(y)|}$  for  $i \in I(y)$  and  $\alpha_i(y) := 0$  otherwise. Then,

$$\|\alpha(y)\|_2 = 1, \quad d_\alpha(x, y) \geq t/\sqrt{|I(y)|} \geq t/\sqrt{r} \quad \forall x \in A$$

- Thus,  $d_T(y, A) \geq t/\sqrt{r}$  for every  $y \in B$

$$\Pr[f \leq r - t] \Pr[f \geq r] \leq \Pr[\mathbf{x} \in A] \Pr[d_T(\mathbf{x}, A) \geq t/\sqrt{r}] \leq e^{-t^2/(4r)}$$

## Theorem (Talagrand's inequality for certifiable functions).

Let  $\mathbf{x} \sim \Omega$  with independent coordinates. Let  $f: \Omega \rightarrow \mathbb{R}$  be 1-Lipschitz with respect to Hamming distance on  $\Omega$ . Suppose that  $\{x \in \Omega : f(x) \geq r\}$  is  $r$ -certifiable. Then, for  $m = \text{med}(f(x))$ ,

$$\Pr[f(\mathbf{x}) \leq m - t] \leq 2e^{-t^2/(4m)}$$

$$\Pr[f(\mathbf{x}) \geq m + t] \leq 2e^{-t^2/(4(m+t))}$$

*Proof.*

$$\Pr[f \leq r - t] \Pr[f \geq r] \leq e^{-t^2/(4r)}$$

- The lower tail in the theorem follows from taking  $r = m$
- The upper tail in the theorem follows from taking  $r = m + t$





# Application 1: Longest Increasing Subsequence

An **increasing subsequence** of a permutation  $\sigma = (\sigma_1, \dots, \sigma_n)$  is defined to be some  $\sigma_{i_1} < \dots < \sigma_{i_\ell}$  for some  $i_1 < \dots < i_\ell$ .

How well is the length  $X$  of the longest increasing subsequence of uniformly random permutation concentrated?

- Talagrand's inequality  $\Rightarrow X = \Theta(\sqrt{n}) \pm \mathcal{O}(n^{1/4})$  w.h.p.
- **Final remark:** *the correct order of the fluctuation is  $n^{1/6}$  (Baik-Deift-Johansson '99). They showed that  $n^{-1/6}(X - 2\sqrt{n})$  converges to the **Tracy–Widom distribution***

# Application 2: Euclidean TSP

Let  $x_1, \dots, x_n \in [0,1]^2$  be uniformly random points in the unit square. The **travelling salesman problem (TSP)** is to find a tour through all the  $n$  points with the shortest possible length:

$$\text{TSP}(x_1, \dots, x_n) := \min_{\pi \in \mathcal{S}_n} \sum_{i=1}^n d(x_{\pi(i)}, x_{\pi(i+1)}), \quad x_{\pi(n+1)} := x_{\pi(1)}$$

Here,  $d(x, y) = \|x - y\|_2$  is the Euclidean distance

- Let  $L_n := \text{TSP}(x_1, \dots, x_n)$  be the random variable of TSP length
- It is known that  $\mathbb{E}[L_n] = \Theta(\sqrt{n})$
- We can show that  $L_n$  is **16-subgaussian**



**Mona Lisa TSP Challenge**

# Application 2: Euclidean TSP

Our plan is to apply the following version of Talagrand's inequality:

**Theorem.** Let  $\mathbf{x} \sim \Omega$  with independent coordinates. Suppose that  $f(y) \geq f(x) - \sum_{i=1}^n \alpha_i(x) \mathbf{1}[x_i \neq y_i]$  for any  $x, y \in \Omega$ . Then, for  $v^2 := 4 \sup_{x \in \Omega} \|\alpha(x)\|_2^2$ ,  $\Pr[|f(\mathbf{x}) - \text{med}(f(\mathbf{x}))| \geq t] \leq 4e^{-t^2/v^2}$

- Let  $\Omega = \{X := (x_1, \dots, x_n) : x_i \in [0,1]^2\}$  and  $f(X) := \text{TSP}(x_1, \dots, x_n)$
- We need to construct a certificate  $\alpha(X)$  such that for any two inputs  $X$  and  $Y$ ,

$$f(X) \leq f(Y) + \sum_{i=1}^n \alpha_i(X) \mathbf{1}[x_i \neq y_i]$$

- We'll show how to merge a tour of  $x$  and the optimal tour of  $y$  and obtain a tour of  $x \cup y$  of length  $\ell_{X \cup Y} \leq f(Y) + \sum_{i=1}^n \alpha_i(X) \mathbf{1}[x_i \neq y_i]$
- Then, by removing the points not in  $x$ , the length is non-increasing. Thus,  $f(X) \leq \ell_{X \cup Y}$

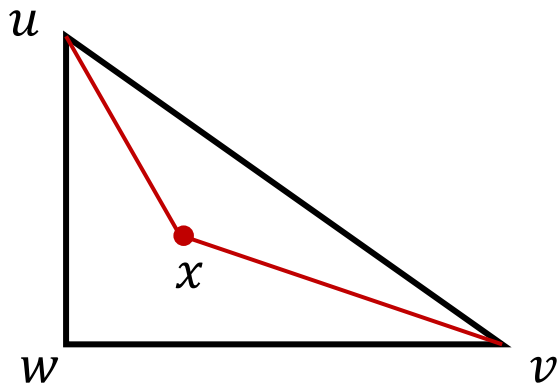
# Application 2: Euclidean TSP

We need the following geometric lemma (related to the Sierpiński curve):

**Lemma.** For any  $x_1, x_2, \dots, x_n \in [0,1]^2$ , there exists a permutation  $\sigma$  such that

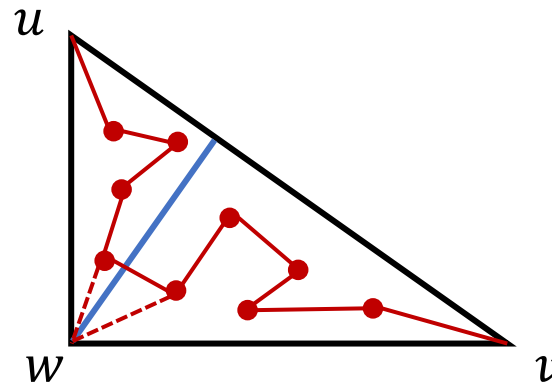
$$d(x_{\sigma(1)}, x_{\sigma(2)})^2 + d(x_{\sigma(2)}, x_{\sigma(3)})^2 + \dots + d(x_{\sigma(n-1)}, x_{\sigma(n)})^2 + d(x_{\sigma(n)}, x_{\sigma(1)})^2 \leq 4$$

*Proof.*

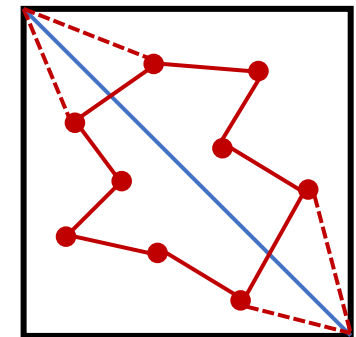


**Pythagorean Inequality:**

$$d(x, u)^2 + d(x, v)^2 \leq d(u, v)^2$$



$$d(u, x_{\tau(1)})^2 + \sum_{i=1}^{m-1} d(x_{\tau(i)}, x_{\tau(i+1)})^2 + d(x_{\tau(m)}, v)^2 \leq d(u, v)^2$$



Merge two triangles

# Application 2: Euclidean TSP

We need the following geometric lemma (related to the Sierpiński curve):

**Lemma.** For any  $x_1, x_2, \dots, x_n \in [0,1]^2$ , there exists a permutation  $\sigma$  such that

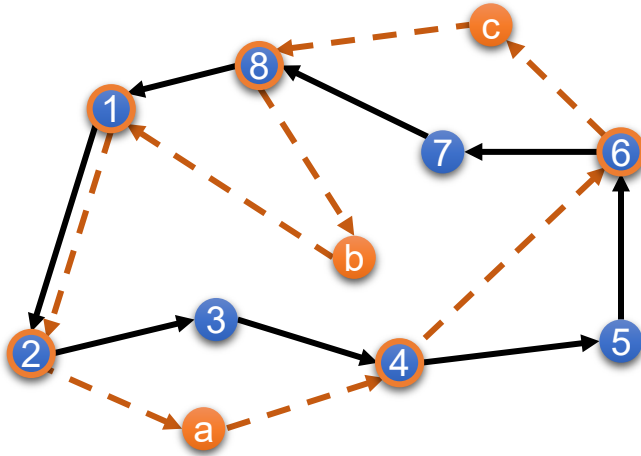
$$d(x_{\sigma(1)}, x_{\sigma(2)})^2 + d(x_{\sigma(2)}, x_{\sigma(3)})^2 + \dots + d(x_{\sigma(n-1)}, x_{\sigma(n)})^2 + d(x_{\sigma(n)}, x_{\sigma(1)})^2 \leq 4$$

- For simplicity, we can consider  $\sigma$  as a function  $\sigma: \mathbb{R}^2 \rightarrow \mathbb{R}^2$  such that  $\sigma(x_{\sigma(i)}) := x_{\sigma(i-1)}$  for any  $i \in [n]$  and  $x_{\sigma(0)} := x_{\sigma(n)}$ , i.e., the **predecessor** function
- Then, the lemma is equivalent to

$$\sum_{i=1}^n d(x_i, \sigma(x_i))^2 \leq 4$$

# Application 2: Euclidean TSP

How to merge the tours

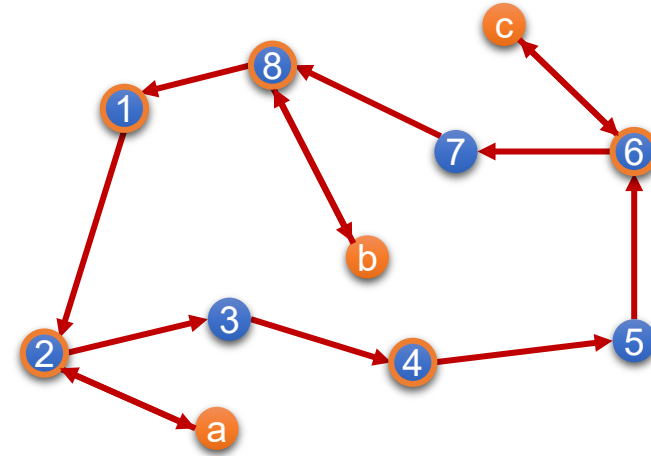


$$y = \{1, 2, 3, 4, 5, 6, 7, 8\}$$

$$f(y) = \text{dist}(1 \rightarrow 2 \rightarrow \dots \rightarrow 8 \rightarrow 1) \text{ (optimal)}$$

$$x = \{1, 2, 4, 6, 8, a, b, c\}$$

Lemma  $\Rightarrow \sigma = \begin{pmatrix} 1 & 2 & 4 & 6 & 8 & a & b & c \\ b & 1 & a & 4 & c & 2 & 8 & 6 \end{pmatrix}$

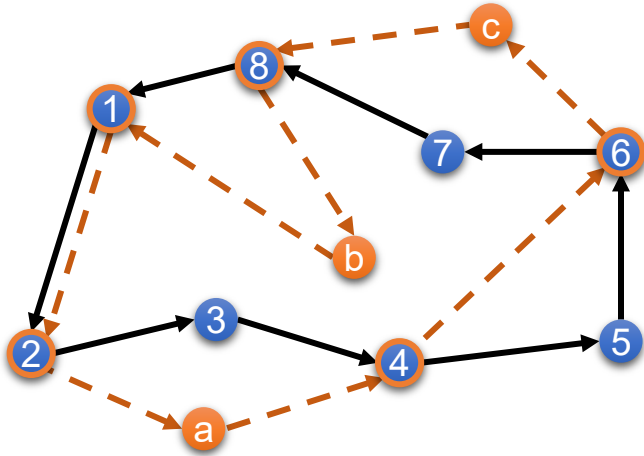


$$1 \rightarrow 2 \rightarrow a \rightarrow 2 \rightarrow 3 \rightarrow 4 \rightarrow 5 \rightarrow 6 \rightarrow c \rightarrow 6 \rightarrow 7 \rightarrow 8 \rightarrow b \rightarrow 8 \rightarrow 1$$

1. Traverses along the optimal order of  $Y$
2. If the current point  $y_k = x_{\sigma(i)}$  and  $x_{\sigma(i+1)} \notin Y$ :
  - i. Traverse along  $X$ 's tour just before it rejoins  $Y$ 's tour
  - ii. Traverse backward and return to  $y_j$

# Application 2: Euclidean TSP

How to merge the tours

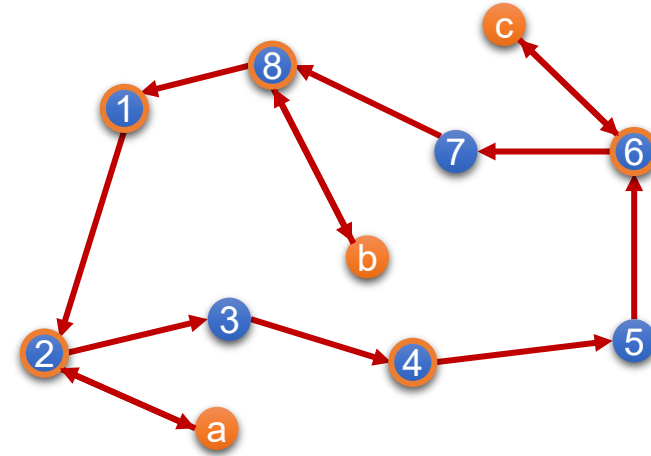


$$y = \{1, 2, 3, 4, 5, 6, 7, 8\}$$

$$f(y) = \text{dist}(1 \rightarrow 2 \rightarrow \dots \rightarrow 8 \rightarrow 1) \text{ (optimal)}$$

$$x = \{1, 2, 4, 6, 8, a, b, c\}$$

Lemma  $\Rightarrow \sigma = \begin{pmatrix} 1 & 2 & 4 & 6 & 8 & a & b & c \\ b & 1 & a & 4 & c & 2 & 8 & 6 \end{pmatrix}$



$$1 \rightarrow 2 \rightarrow a \rightarrow 2 \rightarrow 3 \rightarrow 4 \rightarrow 5 \rightarrow 6 \rightarrow c \rightarrow 6 \rightarrow 7 \rightarrow 8 \rightarrow b \rightarrow 8 \rightarrow 1$$

$$\begin{aligned} f_{X \cup Y} &\leq f(Y) + \sum_{i=1}^n 2d(x_i, \sigma(x_i)) \mathbf{1}[x_i \notin Y] \\ &\leq f(Y) + \sum_{i=1}^n 2d(x_i, \sigma(x_i)) \mathbf{1}[x_i \neq y_i] \end{aligned}$$

# Application 2: Euclidean TSP

$$f(X) \leq f(Y) + \sum_{i=1}^n 2d(x_i, \sigma(x_i)) \mathbf{1}[x_i \neq y_i]$$

- Let  $\alpha_i(X) := 2d(x_i, \sigma(x_i))$
- The **geometric lemma** gives that

$$\|\alpha(X)\|_2^2 = 4 \sum_{i=1}^n d(x_i, \sigma(x_i))^2 \leq 16$$

- Thus, by Talagrand's inequality,

$$\Pr[|f(\mathbf{X}) - \text{med}(f(\mathbf{X}))| \geq t] \leq 4e^{-t^2/16}$$

- That is,  $L_n = f(\mathbf{X})$  is 16-subgaussian

